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Payment Systems Innovation and Currency Demand in India: Some Applied Perspectives

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In this study, we postulate currency demand for transaction purposes driven by income effect, and a payment technology induced substitution effect working through velocity of currency. Innovations in payment systems have shown a statistically significant long-run inverse relationship with currency demand in India. However, the magnitude of its coefficient indicates that the substitution effect on currency demand could be smaller than the dominant income effect. In the presence of payment indicators and financial variables, the income effect coefficient on currency demand has come closer to unity, in line with the standard quantity theory of money but different from the inventory theoretic transaction demand model prediction (0.5). The impact of 1 per cent growth in digital retail transactions volume was estimated to be one tenth of the income effect. Thus, to neutralise the dominant income effect, the payment systems need rapid growth to the extent of 100 per cent in digital retail transactions volume.

JEL Classification: E41, E47, E51, C22

Keywords: Money demand, monetary policy, payment system, error correction and cointegration, ARDL model

Introduction

India's payment systems have witnessed significant growth in digital transactions, supported by policy thrust, sustained efforts from banking and financial sectors to provide technology enabled services and consumers' adoption of various non-cash payment instruments. At the same time, the demand for currency also exhibits growth momentum. Illustratively, currency in circulation to GDP ratio, after declining from 12.1 per cent in 2015-16 to

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8.7 per cent in 2016-17 in the wake of demonetisation¹, recovered to 10.7 per cent in 2017-18 and 11.2 per cent in 2018-19. In this context, a question arises on the impact of payments technology innovation on currency demand. The empirical studies for a cross-section of countries do not offer unique answers. A common perspective is that the relationship of payment systems with currency demand can be country specific. Therefore, we examine this relationship in the Indian context. We analyse currency demand as a function of payment indicators along with other proximate determinants such as income and financial variables (interest rate on deposits and stock return). We are specifically interested in assessing whether payment indicators could impinge on long-run currency demand, for which we use the autoregressive distributed lag (ARDL) model based on quarterly data. The rest of the study is organised into seven sections. Section II provides details of the policy approach to payment systems development in India. Section III reviews the existing literature on payment systems. Section IV gives details of the analytical framework used in this study. Section V discusses data and empirical modelling. Section VI explains stylised facts, Section VII explains empirical findings, and Section VIII concludes the study.

Section II Payment Systems Development: India's Policy Approach

The central bank of a country is usually the driving force behind the development of national payment systems. India's central bank, the Reserve Bank of India (RBI), has been playing this developmental role and has taken several initiatives for safe, secure, sound, efficient, accessible and authorised payment systems in the country (Gupta and Gupta, 2013). A chronology of major initiatives since the late 1980s in the Indian payment systems is given in the Appendix (Table A.1).

The initial steps towards establishing modern payment systems were taken in the early 1980s, when the RBI introduced the magnetic inkcharacter recognition (MICR) technology for cheque processing, which sowed the seeds for digital payment systems in the country. Migration to

¹ The Government of India, on November 8, 2016, announced that ₹500 and ₹1,000 denomination bank notes issued by the RBI shall cease to be legal tender.

the cheque truncation system (CTS) happened in 2008, when the RBI first implemented it in New Delhi, on February 1, 2008 with 10 pilot banks. Soon thereafter the deadline was set to April 30, 2008 for all banks to migrate to the CTS. Further, to handle bulk payments and receipts, the electronic clearing system (ECS) was introduced, which has undergone many changes from being local to regional and then national. For a pan-India system for processing bulk and repetitive payments, the ECS has been subsumed into the National Automated Clearing House (NACH). The next step towards electronic products was the introduction of interoperable ATMs, wherein the National Financial Switch (NFS) had proved that in a large country like India, networked ATMs can function very well.

Payment systems have evolved over time to meet the requirement of remittances using non-cash and non-paper payment methods. The popularly known retail system is the National Electronic Funds Transfer (NEFT). Besides NEFT, the Immediate Payment Service (IMPS) and Real-time Gross Settlement (RTGS) also meet users' funds transfer requirements. Recently, the NEFT has been made operational on a 24x7 basis from December 16, 2019 to ensure availability of digital payments at any time. The IMPS is also a 24x7 immediate funds transfer system. The RTGS is meant for processing large value transactions of above ₹200,000. apart from interbank transactions. The modernisation of the information technology (IT) systems of banks and their core banking solutions has made possible the integration of various delivery channels. The large number of mobile phone users and the availability of cost-effective internet data have led to an increase in the number of mobile internet users. Taking advantage of this, an increasing number of payment facilities are being integrated through the mobile channel. Further, the introduction of a unified payment interface (UPI) has revolutionised the mobile payment system.

A wide range of reforms have been introduced to promote digital payments, covering customer-initiated transactions and government payments. From the perspective of financial inclusion and digitisation of government payments, the use of Aadhaar² for beneficiary identification and authentication in payments has played an important role, enhancing

² Aadhaar is a 12-digit number issued by the Unique Identification Authority of India (UIDAI) to the residents of India based on their demographic and biometric information.

efficiency and transparency in transactions. Accordingly, the Aadhaar Payment Bridge System (APBS) has been put in place to facilitate bulk and repetitive government payments to Aadhaar-seeded bank accounts of identified beneficiaries. Similarly, the Aadhaar Enabled Payment System (AePS) facilitates operations from Aadhaar-seeded bank accounts using biometric authentication of customers. Today, AePS is increasingly being used for Business Correspondent (BC) operations in an interoperable manner. Another significant segment of retail electronic payments is the RuPay cards issued under the Pradhan Mantri Jan Dhan Yojana, with its associated benefits dependent upon usage of the card.

The entry of non-bank players in payment systems has made an important contribution in terms of innovations and convenience to customers. Besides setting up White Label ATMs (WLAs) to bridge the gap in ATM infrastructure in rural and semi-urban areas in the country, nonbanks are actively involved in issuance of Pre-paid Payment Instruments (PPIs). There are many payment systems with significant potential to influence the payment habits of different individuals. The Bharat Bill Payment System (BBPS) provides facility of anytime, anywhere, anyhow bill payments in the country and supports all forms of electronic payments. Also, the Trade Receivables Discounting System (TReDS) caters to the financing needs of the micro, small and medium enterprises (MSMEs) for faster financing and liquidity requirements. Thus, the retail payments ecosystem has made revolutionary progress over the years. The standards of all the payment channels and systems along with their security features in India are comparable with the best in the world. Compared with other countries, the changes in India's payments ecosystem have been fastforwarded to reach its present stage in the shortest possible time. User trust in the payments system is being strengthened through policy initiatives over the years with implementation of certain policies relating to cyber security, fraud, customer awareness, and customer protection. Finally, the RBI has released its payment vision documents in the public domain. According to the latest document, 'Payment and Settlement Systems in India: Vision

³ On May 15, 2019 the RBI announced a vision document titled 'Payment and Settlement Systems in India: Vision 2019–2021', with its core theme of 'Empowering Exceptional (E) payment Experience'. The document aims to empower every Indian with access to a bouquet of e-payment options that are safe, secure, convenient, quick and affordable.

2019-2021³, digital transactions through UPI/IMPS are likely to register average annualised growth of over 100 per cent, which would help in the reduction of currency demand over the vision period.

Section III The Literature

Economics literature on payment systems has evolved in various dimensions, broadly through macroeconomic and microeconomic perspectives (Kahn and Roberds, 2009). The macroeconomic perspective focuses on the impact of payment technology on currency and money demand (Amromin and Chakravorti, 2007; Bech et al., 2018; Columba, 2009; Duca and VanHoose, 2004; Durgun and Timur, 2015; Fischer, 2007; Humphrey, 2004 and 2010; Lippi and Secchi, 2009; Oyelami and Yinusa, 2013). Some studies address the liquidity effects of payment systems (Li and Carroll, 2011), while some others consider the substitution effect of mode of payments (Schuh and Stavins, 2010). On the other hand, microeconomic studies offer perspective on the various determinants affecting the behaviour of people towards the adoption of modern payment instruments and alternate money such as the availability of technology, internet facility and education (Basnet and Donou-Adonsou, 2016; Lippi and Secchi, 2009), card prices (Hazra, 2017; Scholnick et al., 2008), switching and search costs (Ausubel, 1991; Calem and Mester, 1995; Calem et al., 2006; Stango, 2002), interest cost impact on card debt (Brito and Hartley, 1995; Lee, 2014), transaction cost and opportunity cost (Klee, 2008), impact of rewards, discounts and incentives on the use of credit and debit cards (Arango et al., 2015; Arango-Arango et al., 2018; Simon et al., 2010; Stavins, 2018; Valverde and Zegarra, 2011), impact of safety and security aspects (Eze et al., 2008; Kosse, 2013), usefulness of record-keeping of electronic payment data (Galbraith and Tkacz, 2018; Lotz and Zhang, 2016), consumer preference for a specific payment instrument and the impact of price elasticity on instrument choice (Stavins, 2018) and the evidence of beta convergence in the European countries' payment system (Martikainen et al., 2015).

The literature provides mixed evidence on the role of payment systems innovation on money and currency demand. The differential impact is largely due to differences in payment instruments. Illustratively, using disaggregated provincial data in Italy, Columba (2009) shows that innovations in transaction technology through diffusion of ATMs and POS have a negative effect on the demand for currency in circulation but a positive effect on narrow money demand due to the positive effect of technology on bank deposits. Bouhdaoui et al. (2014) analyse the relationship between convenient prices⁴ and cash usage by French nationals and find that individuals' shares of cash payments increased with convenient prices. Clearly, price rigidity can, in part, be explained by the use of cash to pay convenient prices. For Japan, Fujiki and Tanaka (2014), using household-level survey data and a quantile regression model, provide evidence that users of electronic money held more currency than non-users. Bech et al. (2018), using a cross-country (advanced and emerging market economies) panel data model, suggest that despite the increasing use of digital payments across the world, cash continues to be the more preferred mode of payment due to store of value motive and lower opportunity cost rather than payment needs. Some studies also find modern payment instruments to have a positive effect on tax collection – cashless payments reduce tax evasion as the transactions can be tracked easily by the authorities (Hondroyiannis and Papaoikonomou, 2017; Immordino and Russo, 2018).

In the case of developing nations, Oyelami and Yinusa (2013), using data on the Nigerian economy, provide evidence that internet payments and mobile money substituted currency while credit cards, Automated Teller Machine (ATM) and Point of Sale (POS) complimented it. Moreover, barring ATM debit cards and internet payments, all payment channels showed an inverse relationship with shocks to the interest rate and currency demand. This finding may have serious implications for how monetary policy is implemented, especially in developing countries. For India, Nachane *et al.* (2013) find that despite the emergence of various alternatives to cash-based transactions, currency retained its predominance. There exists a cointegrating relationship between currency in circulation, gross domestic product (GDP), wholesale price index (WPI) and deposit rates. The RBI report on the macroeconomic impact of demonetisation (RBI, 2017) analyses growth rates

⁴ Convenient prices are round prices that usually match monetary denominations. Firms set convenient price to avoid cash payments when the cost of handling cash is high or risky (e.g. theft) compared to the costs of other payment instruments. The same reasoning applies to consumers who also face transaction and holding costs when they use cash or other payment instruments (Bouhdaoui *et al.*, 2014).

in various payment channels, observing that an important consequence of demonetisation has been the sharp increase in the use of digital transactions. Maiti (2017) finds that post demonetisation cash transactions have moved in a sustained manner to non-cash modes of payment. Reddy and Kumarasamy (2017) examine the impact of credit and debit cards usage on currency demand in India by employing the ARDL approach. Their results show that the usage of credit cards is negatively associated with currency demand, whereas the usage of debit cards shows a positive association with currency demand in India.

Some studies find that payment technology innovation can also affect income effect on transaction demand for currency and money. In the case of Switzerland, Fischer (2007) re-examines the estimates for income elasticity of money demand based on cross-regional (cantonal) data. The study estimates that income elasticity can range between 0.4 and 0.6. On the contrary, Kumar (2011) analyses data on 20 developing Asian and African countries and finds that the magnitude of income elasticity did not change significantly with the increase in the use of modern payment systems. In the case of India, the income elasticity of currency demand is somewhat higher in comparison to the long-term elasticity observed in similar studies for advanced countries (Nachane *et.al.*, 2013).

Section IV Analytical Approach

From a theoretical perspective, the role of payment systems in currency demand derives from its impact on the velocity of money (Meltzer, 1978). Let us consider the standard quantity theory of money or Fisher's version of the transaction demand for money M (currency with public), driven by velocity of money in circulation (V), economic activity (Y) and general price level (P):

$$MV = PY \tag{1}$$

which can be written in log-linearised statistically estimable form as,

$$LnM_t = \mu + LnY_t + LnP_t + e_t \tag{2}$$

According to this approach, velocity of money V is a constant, thus $\mu = Ln(V)$, for it is driven by technological progress which could be slow

moving. Relaxing this assumption, velocity can change due to rapid progress in payments technology (S) and S_t^j representing different payment indicators, we can rewrite equation (1) as:

$$M V(\mu^{\nu}, S_t^{J}) = PY$$
(3)

Taking log-linearised form,

$$LnM_t = \mu^{\nu} + \theta LnS_t^j + LnY_t + LnP_t + e_t^{\nu}$$
(4)

where $LnV(\mu^{\nu}, S_t^j) = \mu^{\nu} + \theta LnS_t^j$. Thus, in line with the quantity theory of money, it can be shown that

$$LnV(\mu^{\nu},S) = LnM_t - LnP_t - LnY_t = \mu^{\nu} + \theta LnS_t^j + e_t^{\nu}$$
(5)

Empirical studies, however, follow generalised versions of money demand while allowing income elasticity of demand for money deviating from unity (since national income or GDP may not be exactly equal to Fischer's definition of economic transaction) but suppressing technological progress induced change in velocity,

$$LnM_t = \mu + \alpha LnY_t + LnP_t + e_t \tag{6}$$

Or

$$LnM_t - LnP_t = \mu + \alpha LnY_t + e_t \tag{7}$$

It is also evident that some studies estimate nominal demand for money such as

$$LnM_t = \mu + \alpha^n LnY_t^n + e_t \tag{8}$$

Studies considering payments system indicator induced technological progress in velocity would consider

$$LnM_t = \mu^{\nu} + \theta LnS_t^J + \alpha^{\nu}LnY_t + LnP_t + e_t^{\nu}$$
⁽⁹⁾

Or

$$LnM_t - LnP_t = \mu^{\nu} + \theta LnS_t^j + \alpha^{\nu}LnY_t + e_t^{\nu}$$
(10)

Alternatively, for nominal demand for cash,

$$LnM_t = \mu^{\nu,n} + \theta^n LnS_t^j + \alpha^{\nu,n} LnY_t^n + e_t^{\nu,n}$$
(11)

Specific payment indicator S can either increase or decrease money demand depending upon the sign of its coefficient θ . Also, it can induce a

differential income elasticity of demand for money, when $\alpha \neq \alpha^{\nu}$. Similarly, we can have different intercepts $\mu \neq \mu^{\nu}$, reflecting on the scale effect. Furthermore, deriving from Baumol's inventory model of cash demand and Tobin's portfolio and liquidity preference model of money demand, studies incorporate interest rate and asset prices in empirical models of currency and money demand. Finally, a payment indicator can be considered in either volume terms (real transactions) or value (nominal) terms. Both will have different implications on the estimated demand for cash and income elasticity of demand for money.

Section V Data and Empirical Modelling

For the econometric estimation of currency demand in line with the current literature and analytical insights, we use an autoregressive distributed lag (ARDL) model owing to Pesaran and Shin (Pesaran et al., 2001). The choice of the model is partly determined by the sample period for which the data is available. Illustratively, data on currency with the public, part of money demand are available for a very long period from 1950-1951. However, quarterly data on its main determinant, the aggregate income or GDP, are available only from 1996-1997. On the other hand, payments indicators data are available on a monthly basis from April 2004, allowing us to generate quarterly series data for the sample period 2004:2 to 2019:1. Thus, we have a common sample period 2004:2 to 2019:1, *i.e.*, about 60 guarters data, which is not sufficiently large to allow time series models such as vector autoregression (VAR) and vector error correction (VEC) models. Moreover, the advantage of the ARDL model is that it can encompass a mix of stationary and non-stationary variables and allows the estimation of short-run error correction and long-run cointegration relationships among the variables. Since the ARDL model is quite popular, we have avoided offering its technical details.

The ARDL is preferred to other cointegration methods by Engle and Granger (1987), Johansen (1991), and Johansen and Juselius (1990) for various reasons: (i) it provides unbiased estimates of long-run model and valid t-statistics, even when some of the explanatory variables are endogenous (Ali *et al.*, 2016); (ii) it also facilitates short-run analysis: the dynamic error correction model is derived from a simple linear transformation in the ARDL

model; and (iii) being a single equation approach, it can be suitable to a smaller sample. However, the model is not applicable if the order of integration of any of the variables is greater than one, for example, I(2) variable (Menegaki, 2019). In this case, the critical bounds provided by Pesaran et al. (2001) and Narayan (2005) are not valid. The ARDL procedure involves two stages. The first stage is to establish the existence of a long-run relationship. Once such a relationship is established, a two-step procedure is used to estimate the longrun and short-run coefficients of the same equation in the error correction framework. The existence of the long-run relationship is confirmed with the help of an F-test to determine that the coefficients of all explanatory variables are jointly different from zero (Menegaki, 2019). A notable aspect of our empirical exercise is that the payment indicators data exhibit large shifts, reflecting the introduction of new instruments. Large outliers could give rise to biased estimates. In order to overcome this problem, we use the TRAMO/ SEATS approach developed by Gómez and Maravall (1995) to derive smooth linearised series for the payment indicators.

Section VI Stylised Facts

Table 1 provides summary statistics (sample mean, median, standard deviation, *etc.*) of growth rates of currency in circulation, real GDP growth, CPI inflation, and growth of payment indicators in volume and value terms for the period 2004:Q2-2019:Q1. The average growth in payment indicators, both in value and volume terms was two to three times larger than currency growth and four to five times larger than the GDP growth and inflation rate over the sample period.

Table 2 provides summary statistics of nominal and real interest rates and asset returns (equity and gold) which denote the opportunity cost of holding money. During the sample period, real interest rates on bank deposits and medium-term government bond yield (5-year maturity) show a mean of 0.6 percentage points, while assets like equity and gold show higher real returns.

	Table 1:	Summ	ary Stati	stics: Curre	ency, GDP,	Inflation a	nd Paymen	it Indicator	s (Growth	Rates)	
Statistic	GCURP	GY	CPINF	GX50VL	GX53VL	GX54VL	GX55VL	GX50VA	GX53VA	GX54VA	GX55VA
Mean	13.8	7.5	7.2	34.8	34.5	25.7	32.7	23.1	41.9	26.1	28.1
Median	14.9	7.7	6.7	17.9	30.6	24.7	29.1	14.0	37.7	24.4	19.4
Maximum	54.6	13.5	15.1	195.5	81.9	86.5	71.7	122.2	85.1	87.1	219.5
Minimum	-29.1	0.0	2.2	-2.4	3.9	-27.6	-7.3	0.5	20.2	-16.6	0.0
Std. Dev.	11.2	2.5	3.0	52.9	16.1	23.2	19.2	25.6	13.9	21.7	34.0
Skewness	-0.8	-0.6	0.5	2.2	1.0	0.5	0.3	2.6	1.4	0.9	3.9
Kurtosis	9.5	4.0	2.6	6.5	3.8	4.0	2.6	9.3	4.7	4.3	20.4
Jarque-Bera	103.1	5.5	2.4	71.9	10.3	4.8	1.2	153.4	23.5	10.7	831.6
Probability	0.0	0.1	0.3	0.0	0.0	0.1	0.5	0.0	0.0	0.0	0.0
Note: GCURP i to GX55VL rep terms.	s growth rate resent growtl	e of curr h of pay	ency in cird	culation; GY ators in volu	is real GDP me terms; an	growth; CPI d GX50VA t	NF is inflation o GX55VA re	n indicator of epresent grov	consumer pr th rates of p	ice index; wh ayment indice	nile GX50VL ators in value
			Table 2:	Summary	Statistics o	of Interest	Rate and As	sset Return	S		

Statistic	CDINE	CMD	Tau	CBS	DCFD		CMDD	DCFDD	DDTD	CBSD	
DIAUDUC			INIT	200	NHOR			DOLLAN	NINU	NCOD	NULLIN
Mean	6.9	6.5	7.5	7.5	18.2	13.1	-0.4	11.2	0.6	0.6	6.1
Median	6.5	6.5	7.5	7.7	15.6	10.4	0.5	9.6	0.0	1.2	5.4
Maximum	15.1	9.8	9.5	9.0	85.0	52.4	4.1	6.69	4.5	5.0	46.1
Minimum	2.2	3.2	5.2	4.9	-48.5	-12.6	-11.8	-58.7	-8.3	-7.8	-16.7
Std. Dev.	3.0	1.6	1.1	0.8	25.6	15.3	3.4	25.9	2.7	2.8	14.4
Skewness	0.5	-0.2	-0.3	-0.6	0.1	0.5	-1.5	-0.1	-1.2	6.0-	0.6
Kurtosis	2.6	2.6	1.9	3.4	3.8	2.6	5.2	3.6	4.7	3.4	3.0
Jarque-Bera	3.2	1.0	3.9	3.6	1.6	3.0	34.4	1.1	21.3	8.0	3.2
Probability	0.2	0.6	0.1	0.2	0.4	0.2	0.0	0.6	0.0	0.0	0.2

Note: CMR: call money rate; DK1: deposit interest rate; GBS: 5 years Government bond yield; BSER: return on equity; GGOLD: Gold price inflation; CMRR, BSERR, DRTR, GBSR, GGOLDR are indicated real returns of the above variables.

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Section VII Empirical Findings

Since the ARDL method is not applicable when the variables are I(2), we conducted the standard augmented Dickey-Fuller test (ADF) method of unit root test for currency, income and payment system indicators. The unit root test results (Appendix Table A.2) suggest that none of the variables considered in the paper could be characterised with the I(2) process. Hence, the ARDL model could be used for empirical estimation. Since our objective is to examine the long-run impact of payment indicators on currency demand, we begin with the baseline model (M1) that does not include payment indicators and then include alternative payment indicators measured in volume terms in the other models. The different models include different payment indicators, such as retail electronic clearing (REC) including ECS and NEFT (M2), card transactions at POS (M3), retail transactions such as ECS, NEFT, card transactions at POS (M4) and all-digital transactions including BHIM and IMPS (M5). Table 3 provides a summary of the estimated results of the longrun demand function for currency measured in real terms. We derive some interesting perspectives from the estimated cointegrating equations.

First, across the models M2 to M5, coefficients associated with payment indicators have plausible negative sign. Except retail clearing, all other payment indicators have statistically significant coefficients. Thus, in general, we find that the volume growth of non-cash digital payments has a moderating effect on currency demand. Second, the impact differs across payment indicators, as evident from the magnitude of coefficients; illustratively, the coefficient of a narrow measure like POS transactions is twice the coefficient of a broader measure like an all-digital transactions. Third, a crucial finding pertains to the income effect on currency demand, when we compare models with payment indicators (M2 to M5) to models without it (M1). The coefficient of income variable (real GDP) increases when payment indicators are included in the regression, while the intercept term, which captures the deterministic component of velocity, or the scale effect, becomes much smaller. Here, our findings are consistent with the literature (Columba, 2009; Nachane et al., 2013). Fourth, the coefficient of error correction (EC) term associated with the ARDL model, which indicates the speed of adjustment to a long-run path following a short-run deviation, also indicates the role of technology. The EC coefficient is negative and statistically significant, and larger in absolute

Variable	M1	M2 (REC)	M3 (POS)	M4 (retail)	M5 (all-digital)
Real GDP (LY)	0.64 (0.00)	1.36 (0.00)	1.68 (0.00)	1.17 (0.00)	1.35 (0.00)
Payment Indicator (LP)-Volume		-0.17 (0.13)	-0.33 (0.00)	-0.25 (0.00)	-0.18 (0.00)
Intercept	-1.79 (0.03)	-8.32 (0.04)	-10.87 (0.00)	-6.15 (0.00)	-8.08 (0.00)
R-squared	0.94	0.94	0.95	0.96	0.95
SE / SSQ	0.0479 0.1146	0.0479 0.1124	0.0452 0.1000	0.0409 0.0803	0.0464 0.1054
Log Likelihood	93.91	94.46	95.04	103.87	96.26
AIC / SIC	-3.1397 -2.9227	-3.1235 -2.8703	-3.2400 -2.9868	-3.4239 -3.1435	-3.1878 -2.9346
DW-statistic	1.91	1.90	1.93	1.95	1.90
EC	-0.28 (0.00)	-0.30 (0.00)	-0.43 (0.00)	-0.53 (0.00)	-0.45 (0.00)
Null: No residual autocorrelation: AR(2) Test: F-stats (probability) Chi-sq (probability)	1.34(0.27) 2.97(0.23)	1.32(0.28) 2.98(0.22)	1.53(0.22) 3.45(0.18)	0.82(0.44) 1.94(0.38)	1.42(0.24) 3.30(0.19)
F-Bound Test:					
Null: I(0): F stat / 10%, 5% , 1%, critical values	4.71 3.13 3.80 5.38	3.77 2.74 3.29 4.56	5.75 2.74 3.29 4.56	9.48 2.74 3.28 4.56	4.84 2.74 3.29 4.56
Null: I(1): 10%, 5% , 1%, critical values	3.65 4.36 6.03	3.47 4.07 5.59	3.47 4.07 5.59	3.47 4.07 5.59	3.47 4.07 5.59

Table 3: ARDL Model Estimate of Long-run Real Currency Demand Function (Dependent Variable: Natural Log of Currency in Circulation Deflated by CPI)

Note: Figures in parentheses indicate the significance/probability 't' statistic associated with the coefficient.

size for the currency demand equation including payment indicator, than for the equation without it. This implies that payment indicators foster currency demand's adjustment to long-run path through its impact on velocity induced effect. RESERVE BANK OF INDIA OCCASIONAL PAPERS

Table 4 shows the estimates of the long-run demand function for currency in nominal terms. The explanatory variables are nominal GDP and alternative payment indicators measured in volume terms. The results are comparable with the equation estimated for real currency demand, and they provide some notable insights. First, we obtain a statistically significant inverse relationship of payment indicators with currency demand as was the case in the equation for currency demand in real terms. Second, the income

Variable	M1	M2 (REC)	M3 (POS)	M4 (retail)	M5 (all-digital)
Nominal GDP (LYN)	0.92	1.34	1.34	1.14	1.22
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
Payment Indicator		-0.18	-0.24	-0.19	-0.15
(LP)-volume		(0.01)	(0.00)	(0.13)	(0.00)
Intercept	0.05	-3.47	-3.22	-1.52	-2.40
	(0.87)	(0.01)	(0.00)	(0.00)	(0.00)
R-squared	0.99	0.99	0.99	0.99	0.99
SE / SSQ	0.0482	0.0474	0.0454	0.0418	0.0461
	0.1163	0.1104	0.1008	0.0838	0.1043
Log Likelihood	93.49	94.95	97.49	102.68	96.54
AIC / SIC	-3.1248	-3.1409	-3.2317	-3.3816	-3.1980
	-2.9078	-2.8877	-2.9785	-3.0923	-2.9448
DW-statistic	1.93	1.93	1.99	1.90	1.96
EC	-0.30	-0.31	-0.43	-0.43	-0.48
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
Null: No residual	1.27 (0.29)	1.21 (0.31)	2.34(0.11)	0.73(0.49)	1.98(0.15)
autocorrelation:	2.79 (0.25)	2.74 (0.25)	5.07(0.08)	1.71(0.42)	4.35(0.11)
AR(2) Test: F-stats					
(probability) Chi-sq					
F-Bound Test:					
Null: I(0):	7.24	6.26	8.02	11.64	7.35
F stat / 10%, 5% , 1%,	3.13	2.74	2.74	2.74	2.74
critical values	3.80	3.29	3.28	3.29	3.29
	5.38	4.56	4.56	4.56	4.56
Null: I(1):	3.65	3.47	3.47	3.47	3.47
10%, 5%, 1%, critical	4.36	4.07	4.07	4.07	4.07
values	0.03	5.59	5.59	5.59	5.39

 Table 4: ARDL Model Estimate of Long-run Nominal Currency Demand Function

 (Dependent Variable: Natural Log of Currency in Circulation)

Note: Figures in parentheses indicate the significance/probability 't' statistic associated with the coefficient.

elasticity of demand for money without payment indicators is closer to unity but remains above unity when payment indicators are included. Compared with the real currency demand function, nominal income elasticity shows some moderation. Third, the absolute size of the coefficient of payment indicators shows some moderation. Fourth, the size of the intercept term in absolute terms is substantially lower in a nominal demand function.

Next, we examine the impact of payment indicators taken in value terms on currency demand. We estimate both nominal and real currency demand functions. The estimates of nominal currency demand equation are given in Table 5. The retail electronic clearing and POS indicators show a statistically significant inverse relationship with currency demand. On the contrary, the

Variable	M1	M2 (REC)	M3 (POS)	M4 (retail)	M5 (all-digital)
		53	54	50	55
Nominal GDP (LYN)	0.92	1.36	1.34	1.23	0.91
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
Payment Indicator (LP):		-0.13	-0.23	-0.35	-0.05
Value		(0.09)	(0.00)	(0.11)	(0.42)
Intercept	0.05	-3.41	-3.02	0.85	0.65
	(0.87)	(0.10)	(0.00)	(0.03)	(0.03)
R-squared	0.99	0.99	0.99	0.99	0.99
SE / SSQ	0.04	0.04	0.04	0.04	0.04
	0.11	0.11	0.10	0.08	0.09
Log Likelihood	93.49	93.66	95.83	101.99	99.93
AIC / SIC	-3.1248	-3.095	-3.1724	-3.2497	-3.1764
	-2.9078	-2.8420	-2.9192	-2.8518	-2.7785
DW-statistic	1.93	1.92	1.95	2.03	1.97
F-Bound Test:					
Null: I(0):	7.24	5.43	6.85	8.73	6.11
F stat / 10%, 5% , 1%,	3.13	2.74	2.74	2.74	2.74
critical values (Finite	3.80	3.29	3.28	3.29	3.29
sample)	5.38	4.56	4.56	4.56	4.56
Null: I(1):	3.65	3.47	3.47	3.47	3.47
10%, 5% , 1%, critical	4.36	4.07	4.07	4.07	4.07
values	6.03	5.59	5.59	5.59	5.59

Demand Function with Payment Indicators in Value Terms (Dependent variable: Natural Log of Currency in Circulation)

Table 5: ARDL Model Estimate of Long-run Nominal Currency

Note: Figures in parentheses indicate the significance/probability 't' statistic associated with the coefficient.

coefficient of broader measures of payments indicator (all-digital) is negative but statistically not significant.

Results for the real currency demand seem implausible barring the card transactions indicator (Table 6). In the regression equations incorporating all other payment indicators, the coefficient of income is statistically not significant, which seems counter-intuitive. Thus, it is evident that the impact of payments technological innovation on currency demand could be better explained when payment indicators are taken in volume terms rather than in value terms.

Next, to assess the impact of financial variables on currency demand, we estimate the equation by incorporating deposit interest rate and asset

(Dependent Variabl	e: Natural L	og of Curren	cy in Circula	ition Deflate	d by CPI)
Variable	M1	M2 (REC)	M3 (POS)	M4 (retail)	M5 (all-digital)
Real GDP (LY)	0.64 (0.00)	0.48 (0.47)	1.43 (0.00)	0.29 (0.34)	0.15 (0.55)
Payment Indicator (LP)-Value measure		0.03 (0.78)	-0.24 (0.06)	0.16 (0.18)	0.18 (0.02)
Intercept	-1.79 (0.03)	-0.35 (0.95)	-8.46 (0.00)	-0.06 (0.97)	-8.08 (0.00)
R-squared	0.94	0.94	0.95	0.94	0.95
SE / SSQ	0.04 0.11	0.04 0.11	0.04 0.11	0.04 0.11	0.04 0.10
Log Likelihood	93.91	93.93	94.90	94.13	95.41
AIC / SIC	-3.1397 -2.9227	-3.1047 -2.8515	-3.1393 -2.8861	-3.1119 -2.8588	-3.1573 -2.9042
DW- statistic	1.91	1.92	1.90	1.95	1.95
F-Bound Test:					
Null: I(0):	4 71	3 18	4.03	3 50	1 33
F stat / 10% 5%	3 13	2 74	2 74	2 74	2 74
1% critical values	3.80	3 29	3 29	3 28	3 29
1,0,0111001 (01005	5.38	4.56	4.56	4.56	4.56
Null: I(1):	3.65	3.47	3.47	3.47	3.47
10%, 5%, 1%, critical	4.36	4.07	4.07	4.07	4.07
values	6.03	5.59	5.59	5.59	5.59

Function with Payment Indicators in Value Terms (Dependent Variable: Natural Log of Currency in Circulation Deflated by CPI)

Table 6: ARDL Model Estimate of Long-run Real Currency Demand

Note: Figures in parentheses indicate the significance/probability 't' statistic associated with the coefficient.

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(stock) return. The equation for nominal currency demand function did not provide plausible results. The results for the equation of real currency demand function are given in Table 7. Real deposit interest rate and stock

(Dependent variabl	(Dependent variable. Natural Log of Currency in Circulation Denated by CPI)					
Variable	M1	M2 (REC)	M3 (POS)	M4 (retail)	M5 (all-digital)	
Real Income/GDP (LY)	0.64 (0.00	1.10 (0.00)	1.53 (0.00)	1.12 (0.00)	1.16 (0.00)	
Payment Indicator		-0.10	-0.23	-0.23	-0.13	
DRTR*		-0.93	-0.12 (0.74)	(0.01) -0.20 (0.37)	-0.45	
BSER*		-0.09	-0.07 (0.00)	-0.03 (0.12)	-0.06	
Intercept	-1.79 (0.03)	-5.94 (0.00)	-9.54 (0.00)	-5.76 (0.00)	-5.84 (0.00)	
EC		-0.43 (0.00)	-0.51 (0.00)	-0.58 (0.00)	-0.52 (0.00)	
R-squared	0.94	0.94	0.95	0.96	0.95	
SE / SSQ	0.0479 0.1146	0.0474 0.1058	0.0454 0.0967	0.0415 0.0793	0.0467 0.1024	
Log Likelihood	93.91	96.14	98.66	104.22	97.07	
AIC / SIC	-3.1397 -2.9227	-3.1122 -2.7867	-3.2021 -2.8766	-3.3650 -3.0033	-3.1454 -2.8199	
DW-statistic	1.91	1.89	1.92	1.92	1.89	
Null: No residual autocorrelation test: AR(2) Test: F-stats (probability) Chi-sq (probability)	1.34(0.27) 2.97(0.23)	0.79(0.46) 1.90(0.39)	0.96 (0.39) 2.28 (0.32)	0.53 (0.59) 1.31 (0.52)	0.89 (0.42) 2.14 (0.34)	
F-Bound Test:						
Null: I(0): F stat / 10%, 5% , 1%, critical values	4.71 3.13 3.80 5.38	3.05 2.32 2.74 3.71	4.07 2.32 2.74 3.71	6.24 2.32 2.74 3.71	3.42 2.32 2.74 3.71	
Null: I(1): 10%, 5% , 1%, critical values	3.65 4.36 6.03	3.27 3.79 4.97	3.27 3.79 4.97	3.27 3.79 4.97	3.27 3.79 4.97	

 Table 7: ARDL Model Estimate of Long-run Real Currency Demand Function

 (Dependent Variable: Natural Log of Currency in Circulation Deflated by CPI)

Note: Figures in parentheses indicate the significance/probability 't' statistic associated with the coefficient.

*: Coefficients multiplied by 100.

return show negative impact on currency demand, though the interest rate is statistically not significant. Overall, we can say that interest rate and asset returns can marginally influence currency demand. However, the crucial finding here is that with the presence of real interest rate and asset returns, and volume-based payment indicators, barring POS, income effect come closer to unity. Moreover, payment indicators have a statistically significant inverse relationship with currency demand. But the impact is somewhat lower than that found in the equation without financial variable (Table 3), highlighting the role of financial innovation (interest rate and stock return).

For robustness of empirical findings, it is important to reflect on the role of policy to withdraw high denomination specified bank notes (SBNs). Here, we re-estimate two of the equations M1 and M5 from Table 7 for the truncated period up to 2016:Q2; we exclude the period after November 2016 when SBNs were withdrawn from circulation (Table 8). It is crucial to note that we are omitting a small period of 11 quarters, 2016:Q3 to 2019:Q1, which may not be fully adequate to reflect on the policy regime. Nevertheless, this could provide some early reflection. Compared with counterpart equations for the full sample period in Table 7, we find that the income effect is comparable but there is a marginal improvement in the substitution effect of payment indicators on the currency demand. However, the difference in the long-run intercept term, which could reflect upon the scale effect, is somewhat noticeable. Thus, the impact of demonetisation could have impinged in moderating the scale effect on currency demand.

Furthermore, for the stability analysis, we consider the model (M5) of Table 7 and carry out CUSUM and CUSUM square test. The model passed the stability test based on CUSUM test but failed the CUSUM square test, which could be attributed to the period 2016:Q4 for which the model had a relatively larger residual. However, the model could pass the Breusch–Godfrey serial correlation LM test as shown in the estimation tables. In this context, we conclude that since the income effect could be closer to unity, the impact of technology on currency demand could be explored through its impact on velocity of currency in line with equation 5 in Section 4. The suitable ARDL model for velocity of currency, satisfying the stability consideration and no residual serial correlation, could be the one with a structural dummy variable

Variable	M1	M5
Real Income/GDP (LY)	0.76 (0.00)	1.13 (0.00)
Payment Indicator (LP)-volume (all-Digital)		-0.10 (0.01)
DRTR*		-0.27 (0.11)
BSER*		-0.04 (0.00)
Intercept	-2.88 (0.00)	-6.19 (0.00)
R-squared	1.00	1.00
SE / SSQ	0.0075 0.2231	0.0066 0.0016
Log Likelihood	159.17	167.00
AIC / SIC	-6.8074 -6.5665	-7.0222 -6.6610
DW-statistic	1.82	2.18
EC	-0.11 (0.00)	-0.33 (0.00)
Null: No residual autocorrelation test: AR(2) Test: F-stats (probability) Chi-sq (probability)	1.13 (0.33) 2.51 (0.28)	2.18 (0.11) 7.04 (0.06)
F-Bound Test:		
Null: I(0): F stat / 10%, 5%, 1%, critical values	5.30 3.02 3.62 4.94 3.51	5.97 2.20 2.56 3.29 3.09
1 un. 1(1). 1070, 570 , 170, 010000 values	4.16 5.58	3.49 4.37

Table 8: ARDL Model Estimate of Long-run Real Currency Demand Function (Dependent Variable: Natural Log of Currency in Circulation Deflated by CPI;

Note: Figures in parentheses indicate the significance/probability 't' statistic associated with the coefficient.

*: Coefficients multiplied by 100.

for the period since the 2016:Q4 (Table 9 and Appendix Chart A.1). The longrun impact of broader measure of technology on velocity turned out to be statistically significant, though the coefficient size was somewhat lower than the model M5 in Table 7.

Variables	M1	M5
Payment Indicator (LP)-volume	-0.07 (0.00)	-0.08 (0.00)
DRTR*		-0.05 (0.87)
BSER*		-0.06 (0.07)
Structural dummy	0.09 (0.14)	0.10 (0.01)
Intercept	-6.42 (0.00)	-6.39 (0.00)
R-squared	0.99	0.99
SE / SSQ	0.0130 0.0084	0.1221 0.0067
Log Likelihood	170.41	176.92
AIC / SIC	-5.7336 -5.4827	-5.9624 -5.7115
DW-statistic	1.79	1.64
EC	-0.18 (0.00)	-0.27 (0.00)
Null: No residual autocorrelation test: AR(2) Test: F-stats (probability) Chi-sq (probability)	0.37 (0.69) 0.86 (0.65)	0.96 (0.39) 2.44 (0.30)
F-Bound Test:		
Null: I(0): F stat / 10%, 5%, 1%, critical values	11.29 2.63 3.10 4.13	27.67 2.20 2.56 3.29
Null: I(1): 10%, 5% , 1%, critical values	3.35 3.87 5.00	3.09 3.49 4.37

Table 9: ARDL Model Estimate of Long-run	n Velocity of Currency in Circulation
(Dependent Variabl	le: Velocity)

Note: 1. Figures in parentheses indicate the significance/probability 't' statistic associated with the coefficient.

2. Since velocity is defined as the inverse function, V=-(lnM-LP-LY), where LM, LP, and LY refer to natural logarithm of currency in circulation, consumer price index, and real GDP, respectively, the negative coefficient sign in this table may be interpreted with positive impact; increased payment innovations can accentuate velocity and thus, reduce the currency demand.

*: Coefficients multiplied by 100.

Summing up, we found payment technology innovations, as measured in terms of digital volume transactions, having a statistically significant inverse relation with India's currency demand in the long run. However, the substitution effect of digital transactions on currency demand was lower than the strong positive income effect, which suggests that the digital transactions need to increase rapidly if the income effect on currency demand were to be neutralised.

Section VIII Conclusion

The findings of the paper suggest that real income is the major driver of currency demand in India. The income effect of currency demand is closer to unity even when payment indicators, interest rate and asset returns are included as explanatory variables. Payment systems innovation, especially digital transactions in volume terms, showed statistically significant negative effect on currency demand. However, the coefficient of payment indicators was only about one tenth of the income coefficient. Thus, in order to neutralise the real income and inflation induced positive effect on currency demand, digital payments need to increase at a faster pace. Nevertheless, payment systems indicators with growth of around 35 per cent in volume terms during the sample period could have sizeable potential impact of around 3 per cent reduction in currency demand. Digital payments need to grow at a pace of about 100 per cent per annum in order to neutralise the income effect on currency demand. The empirical findings of the study could serve useful for policy purposes. As the payment system grows, and the sample period accumulates with more quarterly data, future studies could exploit the advanced econometric methodology and reassess the impact of technology innovations on currency demand.

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Appendix

Table A.1: Major Milestones in India's Payments and Settlements System'

Year	Major milestone in India's Payment and Settlement Systems
1987	The first ATM machine was introduced by HSBC in Mumbai.
1996	The Institute for Development and Research in Banking Technology (IDRBT) was established by the RBI as an autonomous centre for development and research in banking technology. IDRBT's purpose is to implement a variety of payment applications and foster the development of a reliable communications network. The Governing Council of the IDRBT includes the Deputy Governor and an Executive Director of the RBI, in addition to members from the Indian Banks' Association (IBA) and leading academic institutions (in the areas of science and technology).
1998	The RBI set out its objectives on payment systems in India.
2001	Clearing Corporation of India Limited (CCIL) was setup for clearing & settlement of traders in money, Gsec & foreign exchange markets.
2002	To facilitate faster settlement of trades in government securities in dematerialised form, the RBI introduced an electronic negotiation-based trading and reporting platform called the Negotiated Dealing System (NDS).
_	The subsequent Payment System Vision Document for 2001–04 provided a roadmap for the consolidation, development and integration of the country's payment systems.
2003	The Special Electronic Fund Transfer (SEFT) system was introduced in April 2003 (subsequently discontinued in March 2006 after the introduction of the NEFT).
2004	The Real-time Gross Settlement (RTGS) system introduced to settle interbank and customer transaction above ₹2 lakhs.
2005	National Electronic Funds Transfer (NEFT) system was introduced in November 2005 to facilitate one to one funds transfer of individuals/ businesses.
_	To enhance trading infrastructure in the government securities market, the RBI introduced an electronic order-matching system called the RBI-NDS-GILTS-Order Matching or NDS-OM in short.
_	The Payment and Settlement Systems was detailed in the Vision Document for 2005-08.
_	The RBI constituted the Board for Regulation and Supervision of Payment and Settlement Systems (BPSS) as a committee of its Central Board.

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Year	Major milestone in India's Payment and Settlement Systems
2008	The Payment and Settlement Systems Act, 2007 was enacted empowering the RBI to regulate and oversee all payment and settlement systems in the country and also provide settlement finality and a sound legal basis for netting.
_	The National Electronic Clearing Services (NECS) system, which aims to centralise the Electronic Clearing Service (ECS) operation and bring uniformity and efficiency to the system, was implemented.
_	The National Payments Corporation of India (NPCI), a 'Not for Profit' company, was set up as an umbrella organisation for retail payment systems in India with the guidance and support of the Reserve Bank and the IBA. With initial shareholding of 10 promoter banks, the ownership has since been diversified to 56 banks. The RBI approves the appointment of the Chairman, and the Managing Director and Chief Executive Officer (MD & CEO) of NPCI; it also has placed a nominee director on NPCI's Board. Over the years, NPCI has developed various retail payment products. Taking into account the public sector characteristic of NPCI, the shareholding comprises at least 51 per cent stake by the public sector banks.
_	The Negotiable Instruments Act, 1881 was amended to allow scanned cheque images, paving the way for the cheque truncation initiative that went live in February 2008 in the New Delhi region.
2010	The Vision Document 2009-2012 released on February 16, 2010 reflect the changes after the enactment of the Payment and Settlement Systems Act, 2007, and sets out the objective of ensuring 'that all the payment and settlement systems operating in the country are safe, secure, sound, efficient, accessible and authorised'.
2012	'Payment Systems in India: Vision 2012-15' was announced to proactively encourage electronic payment systems – to ultimately usher in a less-cash society in India – and to ensure that payment and settlement systems in the country are safe, efficient, interoperable, authorised, accessible, inclusive and compliant with international standards.
2016	The 'Payment and Settlement Systems in India: Vision 2018' was announced with the aim of building best in class payment and settlement systems for a 'less-cash' India through responsive regulation, robust infrastructure, effective supervision and customer centricity.
_	In order to achieve the twin objectives of promoting debit card acceptance by a wider set of merchants (especially the small merchants) and ensuring sustainability of the business for the entities involved, the RBI rationalised the Merchant Discount Rate (MDR) framework with effect from January 1, 2018.

Table A.1 (contd.)

Year	Major milestone in India's Payment and Settlement Systems			
2019	The RBI on January 8, 2019 released guidelines on tokenisation for debit/ credit/prepaid card transactions as a part of its continuous endeavour to enhance the safety and security of the payment systems in the country. Tokenisation involves a process in which a unique token masks sensitive card details. Thereafter, in lieu of actual card details, this token is used to perform card transactions in contactless mode at Point of Sale (POS) terminals, Quick Response (QR) code payments, <i>etc.</i>			
_	The RBI released the draft 'Enabling Framework for Regulatory Sandbox on April 18, 2019. Comments on the draft guidelines were invited from stakeholders.			
_	Aiming for a 'cash-lite' society, on May 15, 2019 the RBI announced the 'Payment Systems in India: Vision 2021' to ensure a safe, secure convenient, quick and affordable e-payment system as it expects the number of digital transactions to increase more than four times to ₹8,707 crore in December 2021.			
_	On October 15, 2019 the RBI allowed on-tap authorisation for Bharat Bill Payment Operating Unit, Trade Receivables Discounting System (TReDS) and White Label ATMs.			
_	On September 17, 2019 the RBI released a discussion paper on 'Guidelines for Payment Gateways and Payment Aggregators' and invited feedback from the public.			
_	On September 20, 2019 the RBI announced the harmonisation of turnaround time and customer compensation for failed transactions using authorised payment systems.			
_	On December 16, 2019 the RBI made NEFT available round-the-clock on all days including weekends and holidays.			
_	The RBI introduced a semi-closed Prepaid Payment Instrument up to ₹10,000/- with loading only from a bank account on December 24, 2019.			
2020	On January 1, 2020 the Government of India announced no MDR charges for transactions through RuPay cards and UPI platforms.			
_	The RBI mandated banks not to charge savings bank account customers for online transactions in the NEFT system.			
_	The RBI on January 15, 2020 announced to provide additional security features to card payments by allowing switch on/off and set/modify transaction limits on domestic and international transactions.			

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Sr. No	Variable	Test	Lag Criteria (lags selected)	Test Statistics
1	LCURSA	ADF	AIC (4)	-1.38*
2	LCURRSA	ADF	AIC (4)	-1.30*
3	LYNSA	ADF	AIC (2)	-2.23*
4.	LYS	ADF	AIC (0)	-1.17*
5	LX50VL	ADF	AIC (0)	1.80*
6	LX50VA	ADF	AIC (4)	-1.28*
7	LX53VL	ADF	AIC (4)	-1.31*
8	LX53VA	ADF	AIC (2)	-1.55*
9	LX53VL	ADF	AIC (4)	-1.31*
10	LX53VA	ADF	AIC (2)	-1.55*
11	LX54VL	ADF	AIC (2)	1.09*
12	LX54VA	ADF	AIC (4)	-0.47*
13	LX55VL	ADF	AIC (2)	1.51*
14	LX55VA	ADF	AIC (4)	-0.92*
15	D(LCURPSA)	ADF	AIC (3)	-4.97
16	D(LCURPRSA)	ADF	AIC (3)	-4.96
17	D(LX50VL)	ADF	SIC (0)	-6.69
18	D(LX50VA)	ADF	AIC (3)	-4.35
19	D(LX53VA)	ADF	AIC (1)	-6.93
20	D(LX53VL)	ADF	SIC (0)	-7.81
21	D(LX54VA)	ADF	SIC (2)	-5.20
22	D(LX54VL)	ADF	SIC (1)	-6.68
23	D(LX55VA)	ADF	AIC (4)	-5.19
24	D(LX55VL)	ADF	SIC (0)	-6.50
25	D(LYNSA)	ADF	SIC (0)	-5.56
26	D(LYS)	ADF	SIC (0)	-7.63

 Table A.2: Unit Root Test

Note: *: indicates not significant at 5 per cent (critical value 2.9); thus, cannot reject the null hypothesis that the series has unit root, *i.e.* non-stationary in levels.

Definition:

L: natural log, SA: seasonally adjusted; D: first difference; VA: value, VL: volume, CUR: currency in circulation, CURR: currency in circulation deflated by CPI, YN: nominal GDP, Y: real GDP, X50: retail transactions; X53: retail electronic clearing; X54: card transactions at point-of-sales; X55: all-digital transactions.



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